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Do Divisia monetary aggregates help forecast exchange rates in a negative interest rate environment?

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Abstract

This paper contributes to the literature as the first work of its kind to examine the role and importance of Divisia monetary aggregates and concomitant user cost price indices as superior monetary policy forecasting tools in a negative interest rate environment. We compare the performance of Divisia monetary aggregates with traditional simple-sum aggregates in several theoretical models and in a Bayesian VAR to forecast the exchange rates between the euro, the dollar and yuan renminbi at various horizons using quarterly data. We evaluate their performance against that of a random-walk using two criteria: Root Mean Square Error ratios and the Diebold-Mariano statistic. We find that, under a free-floating exchange regime, superior Divisia monetary aggregates outperform their simple sum counterparts and the benchmark random walk in negative interest rate environment and non-negative interest rate environments, consistently.

Keywords: Forecasting, Exchange rates, Bayesian Vector Autoregression, Uncovered Interest Rate, Sticky Price.

1 Introduction

Forecasting exchange rates is very difficult. Although many economists have written studies on the matter and have found positive results, most of these have later been refuted or at least called into question. There is no one model that works in all circumstances and several authors have argued that none work. In particular, Meese and Rogoff (1983) presented compelling evidence that no model outperforms a driftless random-walk in forecasting exchange rates. Since then, researchers have had a hard time finding a convincing alternative. One such example is Lothian and Wu (2011) which shows that Uncovered Interest Parity (UIP) has remarkable forecasting power in longer time horizons; another is Wright (2008) where the author argues that Bayesian Model Averaging outperforms the random walk in shorter time horizons. Even so, in 2019 Cheung et al. (2019) produced results that reinforced the idea that no model can consistently beat a random walk. None of the aforementioned studies, however, has adopted the approach found in Barnett and Kwag (2006) where they use Divisia monetary aggregates and the User Cost Price within a structural model framework with great success in forecasting the US dollar/British pound exchange rate.

Our objective in this paper is to extend the Barnett and Kwag (2006) experiment by applying it to the Euro/US dollar, US dollar/Yuan, and Euro/Yuan exchange rates (henceforth, EUR/USD, USD/CYN, EUR/CYN) in a negative interest rate environment. In order to do this, we split our data into pre-negative

rates data and the complete data set (which includes negative interest rates). Just as in Barnett and Kwag's study, we employ Divisia Monetary aggregates (Barnett (1978, 1980)) and the User Cost Prices calculated for the Euro zone, the US and China in several structural models and a Bayesian VAR (BVAR) model. In particular, Divisia monetary aggregates replace simple-sum monetary aggregates and the User Cost Price replaces interest rates in each model. We start by evaluating the performance of the Hooper-Morton (HM) model and then proceed to the Flexible Price Monetary model (FP), the Sticky Price (SP) model, Uncovered Interest-rate Parity (UIP), and BVAR. The inclusion of UIP and BVAR in this paper is another innovation with respect to the Barnett and Kwag study. We evaluate the performance of each model using the Root Mean Square Error (RMSE) ratio and the Diebold-Mariano (DM) statistic and compare each model's performance to that of the random walk, as per standard practice in the literature. Each of the aforementioned models becomes HMD, FPD, SPD, and BVARD when it includes Divisia monetary aggregates and the User Cost Price, except UIP which becomes UIPUC, as it does not include Divisia aggregates.

This paper contributes to the literature as the first work of its kind to examine the role and importance of Divisia monetary aggregates and concomitant user cost price indices as superior monetary policy forecasting tools in a negative interest rate environment.

We use quarterly data and the forecasting periods are 1 through 12 quarters ahead. We run the regression for each model twice for each data set: once with the original variables and once with Divisia aggregates and the User Cost Price. We find that, under the RMSE criterion, using Divisia monetary aggregates helps produce forecasts for the EUR/USD that consistently out-perform simple-sum aggregates and the randomwalk in negative and non-negative rates environments using UIP and BVAR; for the EURO/CYN exchange rates, some consistency is observed using the BVAR; no such consistency is found for the USD/CYN. We believe the latter two results have to do with China's foreign exchange policies. Using the DM statistic, results are largely consistent with those under RMSE but are weighed even more heavily toward the use of the more sophisticated Divisia construction for higher levels of forecasting accuracy.

The rest of the paper proceeds as follows: in section 2, we discuss the previous literature related to exchange rate forecasting and Divisia Monetary aggregates; in section 3, we describe what the User Cost Price and Divisia monetary aggregates; in section 4, we refer to the evolution of China's foreign exchange policy; in section 5, we discuss negative interest rates as a policy instrument; in section 6, we present the models; in section 7, we describe the data and their sources present the results, briefly discussing them; section 8 concludes.

2 Literature Review

Forecasting models for exchange rates have existed for decades. PPP and UIP analyses and discussions can be found as far back as the sixties and as recently as 2013 (see, for instance, Balassa (1964) and Lothian and Wu (2011)). Dornbusch (1976) proposed the SP model based on monetary fundamentals and Frankel (1979) further developed this framework by emphasizing the role of expectations. Hooper and Morton (1982) extended this model to include current account balances. But almost immediately after that paper was published, Meese and Rogoff (1983) wrote a seminal study in which they convincingly argued that no exchange rate model can outperform a driftless random walk in out-of-sample forecasting. Since then, Mark (1995) proposed that at longer horizons a monetary fundamentals model could provide better out-of-sample forecasts. This model has been subject to criticism by Faust et al. (2003).

There have been, however, more recent attempts which have shown more promising results: Wright (2008)

the aforementioned Lothian and Wu (2011) are two such cases. Lace et al. (2015) argue that EUR/USD exchange rate can be determined by government yields in the short-run.

BVAR was used in forecasting as far back as Litterman (1986). Sarantis (2006) showed that a BVAR model outperforms a random walk in forecasting daily exchange rates. Banbura et al. (2007) used BVAR for forecasting employment, the Consumer Price Index (CPI) and the Fed Funds Rate with positive results for first-quarter predictions. In a similar fashion, Edge et al. (2010) use several BVAR specifications in order to forecast macroeconomic variables within a DSGE framework, while comparing the accuracy of the forecasts produced by their model to a benchmark model (the FRB/US model). Recently, Schüssler et al. (2018) have used VAR-based models with Bayesian estimation methods for exchange rate forecasting with some success.

More germane to the present study is, of course, Barnett and Kwag (2006) work where they were able to show that the use of Divisia monetary aggregates and the User Cost Price dramatically improve the forecasting power of structural models. In a similar vein, Ghosh and Bhadury (2018) show that Divisia Monetary aggregates are powerful indicators of exchange movements for several economies. The User Cost Price and Divisia Monetary aggregates were derived by Barnett (1978, 1980) which resulted in many volumes of work on monetary aggregation theory and the practical application of these concepts to different areas of economic research. Some of the most important works in the literature (but by no means all of it) has been collected in Barnett and Serletis (2000) and Barnett and Binner (2004). Reimers et al. (2002) found that Divisia aggregates for several countries in Europe have better out-sample-predicting power for the GDP deflator in the Euro area. Similarly, Schunk (2001) showed that using Divisia aggregates improves the accuracy of US real GDP and GDP deflator predictions. Also, Binner et al. (2005) finds there are strong indications that Divisia outperforms simple-sum aggregates in a non-linear framework when forecasting inflation for the euro whilst the predictive power of the user cost spread for economic recessions in both China and the USA has been investigated recently by Chang et al. (2019).

3 Divisia Monetary Aggregates

From the path-breaking work of Barnett (1978, 1980) on microeconomic theory and aggregation theory, we know that the capital stock of money in a given time period is not equal to the monetary service flow (as capital goods do not fully depreciate in a period). The price of these monetary service flows is the opportunity cost, or user cost, of holding a particular monetary asset for that period. The User Cost Price then is the present value of however much interest an agent is foregoing because they are holding an asset, given that there exists a pure investment asset which provides a higher return and no monetary services. The User Cost Price is calculated thus:

$$\pi_{it} = (R_t - \gamma_{it})/(1 + R_t)$$
(1)

where γ_{it} is the return on asset *i* and R_t return on the pure investment or benchmark asset. A key feature of the User Cost Price is that it can never be negative and this is particularly relevant for our results (see section 7).

With the User Cost Price precisely defined, an aggregate for the monetary service flows can be elaborated which will track these flows correctly. For this purpose a Divisia index is used. For the construction of Divisia indexes, let the share weight for each individual asset i over time, t, s_{it} , be defined as

$$s_{it} = \pi_{it} m_{it} / \sum \pi_{jt} m_{jt} \tag{2}$$

where m_{it} is the nominal monetary asset *i* at time *t* and s_{it} is defined as $s_{it} = 1/2(s_{it} + s_{it-1})$. And so, the Divisia monetary index is

$$\ln M_t - \ln M_{t-1} = \sum_{t=1}^n s_{it} (\ln m_{it} - \ln m_{it-1})$$
(3)

Here M_t is the quantity index. From the above equation, one can see that the growth rate of the index is a weighted sum each monetary asset *i*. Each *i* has a share in the User Cost and this is precisely its corresponding weight in the Divisia index. Finally, the accompanying User Cost Price index Π is defined as

$$\ln \Pi_t - \ln \Pi_{t-1} = \sum_{t=1}^n s_{it} (\ln \pi_{it} - \ln \pi_{it-1})$$
(4)

The idea here is that agents substitute toward holding the monetary assets which have the lowest relative user costs whenever there is a change in the own interest rate of another component monetary asset. This reflects how agents take into account opportunity costs in their decision process. The Divisia monetary aggregates and associated user cost price indices internalize the liquidity preferences of the asset holders in the construction of the index via the share weights, s_{it} , of the assets held.

4 China's Financial Markets

Financial markets are eager for any signal of monetary policy from the People's Bank of China (PBC) and the importance of effective monetary policy communication will only increase as China continues to liberalize its financial system and open its economy. The implementation of China's "Open-door policy" has achieved rapid economic growth for three decades, please see Bohnet et al. (1993) for details. The capital inflow through foreign direct investment, together with an abundance of cheap labor together helped China and the whole world enjoy low price goods for over twenty years. Prior to 1994, China applied a dual-core pegged foreign exchange rate domestically and internationally in order to protect its fragile financial system. Since 1994, the Chinese yuan has operated with a currency peg in order to keep its value low compared to other countries. The effect on trade is that Chinese exports are cheaper and, therefore, more attractive when compared to those of other nations. This policy encourages the global marketplace to buy its goods to ensure economic prosperity.

More recently, China's exchange rate regime has undergone gradual reform. After announcing the move away from a fixed exchange rate in July 2005, China began taking regular steps towards a more flexible currency, while exchange rate stability continued to play an important role. The PBC announced that China was "moving into a managed floating exchange rate regime based on market supply and demand with reference to a basket of currencies." The basket of currencies was not specified, however, and the regime in operation was one with a continued tight link to the U.S. dollar. Specifically, there would be a daily rate (the central parity rate, or the 'fix') announced before the start of the trading day that would form the midpoint of the band within which the CYN/USD rate could fluctuate on that day. The renminibi has therefore become more flexible over time but is still carefully managed, and depth and liquidity in the onshore FX market is relatively low compared to other countries with de jure floating currencies. Allowing a greater role for market forces within the existing regime by making central parity formation for the daily trading band (the 'fix') mechanical and transparent is critical for greater two-way flexibility of the exchange rate. The use of FX intervention should be guided by the need to limit excessive volatility; and capital flow management measures (CFMs) should not be modulated to help manage the exchange rate.

Going forwards, further steps to develop the FX market, improve FX risk management, and the development of an alternate monetary policy anchor by continuing to modernize the monetary policy framework are recommended. An overview of the evolution of China's exchange rate regime from 2005 onward, including details of the unique constraints faced by China on its path to a floating exchange rate is provided by Das (2019). China's unique institutional setup and the impact of the PBC's main communication channels on financial markets is provided by McMahon et al. (2018). This detailed analysis of China's monetary policy framework recommends that providing timely information in one place (in Chinese and English), expanding PBC forecasting resources and capacity, and holding regular press conferences would not only be helpful for monetary policy, but also increase the attractiveness of China's capital markets and advance renminbi internationalization.

5 Negative Interest Rates

Negative interest rate policy (NIRP) has become a standard instrument in the ECB's toolkit over time but remains controversial, both in central banking circles and academia. Central banks impose the drastic measure when they fear their national economies are slipping into a deflationary spiral, in which there is no spending, and hence, dropping prices, no profits, and no growth. Most central banks that adopted NIRP were primarily motivated by the stabilization of inflation expectations as NIRP aims to increase the supply of credit by taxing banks' excess reserves at the central bank and thereby support growth, Jobst and Lin (2016). NIRP complements asset purchases and forward guidance that has been implemented since the Global Financial Crisis to ensure that the economy is sufficiently stimulated. In spite of these positive effects on the operation of monetary policy, NIRP has often been criticized for its potential side effects, particularly on the banking sector. A theoretical model that explains how policy rates transmit to banks' supply of credit, is provided by Bittner et al. (2020). A useful summary of existing work on the impact of negative rates on banks' lending and securities portfolios, and the consequences for the real economy are provided in Heider et al. (2021).

Sweden's central bank was the first to deploy negative interest rates in July 2009 when the Riksbank cut its overnight deposit rate to -0.25%. The European Central Bank (ECB) followed suit in June 2014 when it lowered its deposit rate to -0.1%. As experience with negative interest rates was scant, the ECB proceeded cautiously over time, lowering the deposit facility rate (DFR) in small increments of 10 basis points, until it reached -0.5% in September 2019. The ECB turned to negative interest rates to lower the value of the euro. Low or negative yields on European debt will deter foreign investors, thus weakening demand for the euro. Empirical evidence regarding the impact of NIRP on exchange rates is scant, although a survey on recent developments in the monetary policy transmission mechanism in NIRP adopted countries by Ball et al. (2016) concludes that exchange rate appreciation pressures are generally reduced and that the policy has been associated with an improvement in overall financial conditions along with a modest expansion of credit in the euro area. Arteta et al. (2016) suggest that the impact of NIRP on exchange rates has been more varied with currencies depreciating on average against the U.S. dollar and on trade-weighted-terms, except for the Japanese yen and the Swiss franc.

The theoretical challenge is to integrate the role of liquid assets into a model of bank lending. Holding liquid assets and trading them in interbank markets is essential for lending because it allows banks to measure

and manage asset liquidity. As stated in Section 3 above, a sophisticated Divisia index measure, under fairly general assumptions, represents the ideal aggregate measure of "liquidity services" available in the economy and is therefore potentially of great interest to monetary policy-makers aiming at understanding the effects of monetary policy on the aggregate economy, Keating et al. (2019). We follow Chang et al. (2019) and take the view that Divisia method of pricing incorporates the segmented markets hypothesis by treating assets of different degrees of liquidity and different maturities as imperfect substitutes. Indeed, one of the main contributions of the Divisia monetary aggregate literature is to uncover and acknowledge the failings of the simple sum approach that treats all monetary assets as perfect substitutes. Furthermore, the expectations of the interest rates in this case can be put aside as all constituent component assets are treated as imperfect substitutes based on their liquidity and store of value pricing at the present time period, although a more complex expectations hypothesis could be developed as demonstrated in the monetary asset case in Barnett and Wu (2005). The gains in predictive accuracy by incorporating a user cost index in macro forecasting models, including exchange rate forecasting models, is clear; user cost indices provide a unique interpretation for the link between the yield spread and recessions

6 The Models

Hooper and Morton (1982) developed an exchange rate forecasting model which was based on previous models such as the Dornbusch (1976) Sticky Price model and the Flexible Price Monetary model by Frenkel (1976). The HM model includes the Current Account (CA) as an explanatory variable (its principal innovation). Thus, we have the following,

$$e_{t} = \beta_{0} + \beta_{1}(m_{t} - m_{t}^{*}) + \beta_{2}(y_{t} - y_{t}^{*}) + \beta_{3}(i_{t} - it^{*}) + \beta_{4}(p_{t} - p_{t}^{*}) + \beta_{5}ca_{t} + \beta_{6}cat^{*} + \nu_{t}$$
(5)

where e_t is the exchange rate and m_t and m_t^* , y_t and y_t^* , i_t and i_t^* , and p_t , p_t^* , ca_t and ca_t^* are, respectively, domestic and foreign money supply, domestic and foreign output, domestic and foreign interest rates, domestic and foreign current long-run expected rates of inflation, and domestic and foreign current account balances at time t.

The model specification involves an error-correction restriction and so as to avoid short-run dynamics. What this means is that the variation from the exchange rate is a correction of the deviation from a long-run equilibrium in the previous period. Taking the natural logarithms of all variables except the current account variable, the equation becomes the following,

$$lne_{t+h} - lne_t = \alpha_0 + \alpha_1 (lne_t - \beta_0) - \beta_1 ln\tilde{m}_t - \beta_2 ln\tilde{y}_t - \beta_3 ln\tilde{i}_t - \beta_4 ln\tilde{p}_t - \beta_5 ca_t - \beta_6 cat^*) + \epsilon_t$$
(6)

Here \tilde{m}_t , \tilde{y}_t , \tilde{i}_t , and \tilde{p}_t are domestic to foreign relative money supply, output and short-term interest rates, respectively, and h is the forecasting horizon. We should note that we have replaced long-run expected rates of inflation with the only proxy available, relative prices. We will explain this choice in the next section.

Notice that by setting $\beta_5 = \beta_6 = 0$, the model is reduced to the Sticky Price model; $\beta_4 = \beta_5 = \beta_6 = 0$ results in the Flexible Price Monetary model; and, $\beta_1 = \beta_2 = \beta_4 = \beta_5 = \beta_6 = 0$ is Uncovered Interest-Rate Parity.

The BVAR model with a Minnesota prior was introduced in the aforementioned Litterman (1986) and, as previously described, has been widely used in forecasting. If the model is as follows

$$y = (I_m \otimes X)\alpha + \epsilon, \qquad \epsilon \sim (0, \Sigma_\epsilon \otimes I_T)$$
(7)

then y and ϵ are $mT \times 1$ vectors of dependent variables and errors, respectively, and where m is the number of variables and T, the time periods. I_m is the identity matrix, X is the matrix of independent variables and α is a $ml \times 1$ vector where l is the number of lags. More specifically, $\alpha = \bar{\alpha} + \xi_{\alpha}$ with $\xi_{\alpha} \sim N(0, \Sigma_{\alpha})$, where in the Minnesota prior $\bar{\alpha} = 0$ except $\bar{\alpha}_{1i} = 1, i = 1, ..., m, \Sigma_{\alpha}$ is diagonal and each element $\sigma_{ij,l}$ (equation i, variable j, and lag l) is as follows

$$\sigma_{ij,l} = \phi_0/h(l), \qquad i = j \tag{8}$$

If j is endogenous, then

$$\sigma_{ij,l} = \phi_0 \times \phi_1 / h(l) \times (\sigma_j / \sigma_i)^2, \qquad i \neq j$$
(9)

And if j is exogenous, then

$$\sigma_{ij,l} = \phi_0 \times \phi_2 \tag{10}$$

In this case $\phi_0, \phi_1, \phi_2, (\sigma_j/\sigma_i)^2$ and h(l) are, respectively, hyperparameters, a scaling factor, and a function of lags l. Note that ϕ_0 measures the tightness of the first lag's variance, ϕ_1 is the relative tightness of any other variables, and ϕ_2 is the relative tightness of exogenous variables. Finally, h(l) is a measure of the relative tightness of the variance of the lags.

The error correction model follows a similar process to the one laid out for the SP model, using the same variables. The number of lags is 5 for the EUR/USD and USD/CNY and 6 from the EUR/CNY, the averages of three information criteria.

Every one of the above models will be estimated twice: once with their standard variables, and a second time with M3 monetary aggregates replaced by the Divisia index and the reference interest rate replaced by the User Cost Price. Here, the use of the User Cost instead of the interest rate follows Barnett et al. (1984). There are a total of ten models whose forecasting performance will be evaluated. All data are in logs, except interest rates and the User Cost Prices.

6.1 Performance Evaluation

In this study we use a rolling regression in order to produce the predicted forecasts. We first pick an in-sample period for which the models are first estimated and then exchange rates are forecast for the outof-sample period. The sample is then updated to the following period until there are no more out-of-sample observations. In order to pick the in-sample and out-of-sample periods for the whole sample (including negative rates), we chose the date at which interest rates become negative, i.e. June 2014, for the exchange rates involving the euro. For the USD/CYN, we picked January, 2015 as the start of the out-of-sample period, as that signified the end of the 2005-2015 period of exchange rate regime reform. For the pre-negative rates data, the out-of-sample period begins after the end of the Great Recession.

The performance of each model is evaluated by comparing each one to a benchmark model which in this case is the driftless random-walk.

For the first evaluation method we use the root mean square error (RMSE) of each of the models and divide it by the RMSE of the random-walk. A ratio of less than one indicates that the model is performing better than the random-walk and vice-versa.

The second method is the statistic produced by Diebold and Mariano (1995), which allows for the comparison of forecasts in terms of whether the difference between two forecasts for the same forecasting period is statistically significant and whether or not the improvement is statistically significant (one forecast being "better" than another).

7 Data and Results

7.1 Data

The data we utilize are quarterly series of the different variables in the models from 2002Q1 to 2018Q4. We obtained Divisia M3 monetary aggregates and user cost prices for the Euro Area (including the first 12 member countries), US, and China from the Bruegel Institute¹, Center for Financial Stability² and The Center for Financial Development and Stability³ websites, respectively. In terms of our independent variables, we use 3-month Treasury bill rates for the short-term interest rates, quarterly GDP for output, CPI as the price level, and the current account balances. We use traditional M3 monetary aggregates to compare with our theoretically superior Divisia aggregates. All of these were retrieved from the Federal Reserve Economic Data bank found in the Federal Reserve Bank of St. Louis' website⁴.

7.2 Results

Figures 1 and 2 in the annex show the forecasts of the BVARD and UIPUC models against the random walk forecasts and the actual exchange rates for the out of sample period 3, 6, 9, 12 months ahead. Tables 1 and 2 display the RMSE ratios of the different models under consideration without Divisia and with Divisia and the User Cost Price for every forecasting period, the full sample and the sample without negative rates, respectively. What we notice is that, when we compare both tables, the only two models that consistently beat the random walk are the UIPUC and BVARD for the USD/EUR exchange rate. That is to say, all the other models' improvements are either sensitive to in-sample/out-of-sample period changes or show no improvements at all. The one exception appears to be BVARD for the EUR/CNY where the model with Divisia and the User Cost Price beats the random walk with and without negative rates, especially in the short-run. The more moderate forecasting results for USD/CYN and EUR/CYN can be explained by the "managed foreign exchange rate" regime for the Chinese RMB through monetary policy during the foreign exchange market regime reform. In other words, the improvements over the random walk appear only in the context of a floating exchange rate regime. It is also worth noting that UIPUC performs so well because the difference in this model is not between two reference rates but between a weighted basket of returns. In the case of UIP with negative rates, the difference will actually become a sum (of two positive or negative numbers) – the opposite of what happens when we use the User Cost Price which can never be negative by construction. Another point we should mention is that the BVAR with Divisia is the model that behaves as one would expect, considering the literature on Bayesian methods: it has strong forecasting power in the short-run and becomes weaker as we move towards the long-run.

 $^{{}^{1}}https://www.bruegel.org/publications/datasets/divisia-monetary-aggregates-for-the-euro-area/agg$

²http://www.centerforfinancialstability.org/

 $^{^{3}} http://cfds.henuecon.education/index.php/data/chinese-divisa-data$

⁴fred.stlouisfed.org

The DM statistic in Tables 3 and 4 provide supporting evidence for the results found under the RMSE criterion. When comparing the forecasts produced by the models and those produced by the random walk, all models behave similarly to how they performed under the RMSE criterion. For UIPUC for the EUR/USD in the longer forecasting horizons, the DM statistic becomes negative and increasingly so. P-values usually decrease in every period and by the 2-year horizon, the p-value approaches or reaches the 10% significance level. At the 2 and 3-year horizons it is usually below that threshold. The opposite happens with the BVARD for the USD/EUR and the EUR/CNY. This implies that, again, in these cases, models which include the User Cost Price and Divisia aggregates produce forecasts that provide statistically significant improvements on the random walk forecasts and the standard models' forecasts. Also, as mentioned in the previous section, these improvements can only be observed in the exchange rates affected by negative interest rates in the context of free floating regimes. Otherwise, models (with or without Divisia) do not behave consistently or under-perform consistently.

8 Conclusion

This paper is based on solid theoretical foundations and contributes to the literature as the first work of its kind to examine the role and importance of Divisia monetary aggregates and concomitant user cost price indices as superior monetary policy forecasting tools in a negative interest rate environment. We echo Belongia (2006) that the use of user cost price duals appear to be worthy of further investigation. In particular, the sensitivity of inference to changes in measurement alone goes to the core of empirical monetary research. Conventional practice in empirical work and policy discussions have been to knowingly use index numbers that cannot possibly be meaningful representations of either the aggregate quantity of money or its price. Results presented here provide the first available evidence that Divisia monetary aggregates and their concomitant user cost price indices provide superior information about future forecasts of exchange rates in a negative interest rate environment. This finding also holds for the US China exchange rate, which is managed centrally. This result is important for monetary policymakers and academic researchers around the world, particularly given the recent trend towards monetary policymakers' decision to operate in a negative interest rate environment. The Divisia monetary aggregates and associated user cost price indices internalize the liquidity preferences of the asset holders in the construction of the index via the share weights of the assets held. A final inference to draw is that resources directed towards the construction and dissemination of monetary statistics that meet the same standards applied to other economic aggregates are likely to yield a high return in our understanding of exchange rate forecasting in particular and economic activity more generally. Future work will consider further innovations in the construction of the Divisia index to incorporate more sophisticated measures of the riskiness of the assets, building upon Binner et al. (2018) while further work to understand the information channel of monetary policy following monetary policy shocks is recommended Hoesch et al. (2020).

	Pane	A. Hooper-Mortor	a Model RMSE ove	r Random Walk RN	ASE	
	Quarterly EU	R/USD Ratio	Quarterly EU	R/CNY Ratio	Quarterly US	D/CNY Ratio
Time Horizon	HM	HMD	HM	HMD	HM	HMD
1 quarter	1.15478	1.18216	1.05538	1.07073	0.98306	1.00485
2 quarters	1.18824	1.21720	1.08965	1.11423	0.98548	1.02506
3 quarters	1.21131	1.21706	1.06486	1.09073	0.98856	1.06166
4 quarters	1.21403	1.22721	1.02656	1.08095	0.99294	1.11573
6 quarters	1.27698	1.30526	0.98503	1.00352	1.07743	1.26496
8 quarters	1.27922	1.37360	0.97595	0.97220	1.18930	1.47110
12 quarters	1.01244	1.05233	0.70653	0.79511	1.61649	1.86223
	Panel B.	. Sticky Price Mone	tary Model RMSE	over Random Walk	: RMSE	
	Quarterly EU	R/USD Ratio	Quarterly EU	R/CNY Ratio	Quarterly US	D/CNY Ratio
Time Horizon	SP	SPD	SP	SPD	SP	SPD
1 quarter	1.14907	1.20118	1.02008	1.08749	1.02692	1.01659
2 quarters	1.18519	1.25805	1.00275	1.14261	1.04491	1.03122
3 quarters	1.20969	1.29250	0.92323	1.08395	1.05755	1.04812
4 quarters	1.20331	1.33369	0.86323	1.03652	1.07329	1.10239
6 quarters	1.27322	1.41371	0.83207	0.89969	1.17687	1.25307
8 quarters	1.31429	1.50507	0.74822	0.74303	1.35766	1.50606
12 quarters	1.01219	1.08908	0.65601	0.49146	2.13884	2.12648
	Panel C.	Flexible Price Mone	etary Model RMSE	over Random Wal	k RMSE	
	Quarterly EU	R/USD Ratio	Quarterly EU	R/CNY Ratio	Quarterly US	D/CNY Ratio
Time Horizon	FP	FPD	FP	FPD	FP	FPD
1 quarter	1.08955	1.14115	1.05311	1.13516	1.02327	1.01160
2 quarters	1.10019	1.18503	1.03972	1.18799	1.04906	1.03327
3 quarters	1.09983	1.18284	0.93651	1.09647	1.07286	1.06526
4 quarters	1.07143	1.19878	0.86541	1.03314	1.10116	1.13338
6 quarters	1.11431	1.21480	0.82233	0.90563	1.22389	1.29360
8 quarters	1.12991	1.25118	0.67344	0.71554	1.42539	1.54135
12 quarters	0.85684	0.96771	0.55671	0.41385	2.13406	2.09900

Table 1: Annex. RMSE Ratios (Full Sample)

el D. Uncovered Interest Parity Model RMSE over Random Walk RMSE	rly EUR/USD Ratio Quarterly EUR/CNY Ratio Quarterly USD/CNY Ratio	UIPUC UIPUC UIPUC UIPUC UIPUC	0.99558 1.09506 1.06225 1.04892 1.04516	$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	0.91933 1.14697 1.11713 1.16410 1.09721	0.90676 1.12440 1.09801 1.23103 1.12527	0.89967 0.94819 0.95994 1.33835 1.15844	0.82799 0.83637 0.85267 1.44102 1.19331	0.64757 0.52655 0.70845 1.61199 1.18844	nel E. Bayesian Vector Autoregression RMSE over Random Walk RMSE	rly EUR/USD Ratio Quarterly EUR/CNY Ratio Quarterly USD/CNY Ratio	BVARD BVAR BVARD BVAR BVARD BVARD	0.54916 1.08918 0.98758 1.03015 1.04204	0.93742 1.12824 0.99009 1.04885 1.05682	0.98521 1.08190 0.98996 1.05880 1.06096 1.06096 1.05880 1.06096		1.03732 1.10665 0.75972 0.74646 0.82516	1.02910 0.98249 1.03850 1.05651 1.20615	1 01812 1 16328 1 09868 0 83834 1 86274
Panel D. Uncovered Inter	Quarterly EUR/USD Ratio	UIP UIPUC	1.06473 0.99558	1.06618 0.95398	1.04716 0.91933	1.03291 0.90676	1.03354 0.89967	1.03188 0.82799	0.80787 0.64757	Panel E. Bayesian Vecto	Quarterly EUR/USD Ratio	BVAR BVARD	1.04407 0.54916	1.07401 0.93742	1.05377 0.98521	1.01364 1.02364	0.92489 1.03732	1.01270 1.02910	1.00774 1.01812
		Time Horizon	1 quarter	2 quarters	3 quarters	4 quarters	6 quarters	8 quarters	12 quarters			Time Horizon	1 quarter	2 quarters	3 quarters	4 quarters	6 quarters	8 quarters	12 quarters

:	(Continued)
	KMSE Katios
	Annex.
Ē	Table 1:

	Pane	A. Hooper-Mortor	n Model RMSE over	r Random Walk RN	MSE	
	Quarterly EU	R/USD Ratio	Quarterly EUI	R/CNY Ratio	Quarterly USI	O/CNY Ratio
Time Horizon	HM	HMD	HM	HMD	HM	HMD
1 quarter	0.92434	0.99926	1.12076	1.12767	0.72794	0.74062
2 quarters	0.80474	1.05181	1.19166	1.29742	0.60545	0.64119
3 quarters	0.72752	1.11086	1.23004	1.38577	0.54704	0.61586
4 quarters	0.61803	1.11239	1.31869	1.52593	0.44238	0.53150
6 quarters	0.69631	1.59419	1.41764	1.71139	0.31140	0.42953
8 quarters	0.72167	1.86105	1.51426	1.78603	0.21391	0.30494
12 quarters	1.08860	1.84508	1.23119	1.32586	0.14411	0.20044
	Panel B.	. Sticky Price Mone	tary Model RMSE	over Random Walk	: RMSE	
	Quarterly EU	R/USD Ratio	Quarterly EUI	R/CNY Ratio	Quarterly USI	D/CNY Ratio
Time Horizon	SP	SPD	SP	SPD	SP	SPD
1 quarter	0.92625	1.06009	1.23298	1.26597	0.67663	0.68077
2 quarters	0.83837	1.14833	1.39162	1.48124	0.56486	0.57200
3 quarters	0.79795	1.23481	1.43166	1.54206	0.49940	0.50730
4 quarters	0.65749	1.25684	1.49251	1.68579	0.41834	0.43688
6 quarters	0.72019	1.83673	1.60080	1.86142	0.29856	0.34992
8 quarters	0.83810	2.07640	1.63473	1.86210	0.21101	0.25924
12 quarters	1.24603	1.98105	1.29748	1.37265	0.08151	0.21832
	Panel C.	Flexible Price Mone	etary Model RMSE	over Random Wal	k RMSE	
	Quarterly EU	R/USD Ratio	Quarterly EUI	R/CNY Ratio	Quarterly USI	D/CNY Ratio
Time Horizon	FP	FPD	FP	FPD	FP	FPD
1 quarter	0.91977	1.12106	1.19491	1.21359	0.70104	0.69328
2 quarters	0.81436	1.25485	1.36808	1.45019	0.57844	0.56899
3 quarters	0.77264	1.39750	1.45161	1.55190	0.49747	0.49457
4 quarters	0.59926	1.47436	1.51992	1.70846	0.40562	0.39266
6 quarters	0.64261	2.26233	1.61788	1.88977	0.26236	0.26143
8 quarters	0.87746	2.58602	1.64600	1.87995	0.15763	0.19674
12 quarters	1.24750	2.45541	1.30103	1.39501	0.21565	0.27810

Table 2: Annex. RMSE Ratios (W/out Negative Rates)

	USD/CNY Ratio	UIPUC	0.77262	0.63739	0.54518	0.45514	0.34370	0.27096	0.12604		USD/CNY Ratio	BVARD	0.79645	0.68023	0.63755	0.60446	0.38043	0.22741	0.11579
Valk RMSE	Quarterly 1	UIP	0.89740	0.83043	0.78200	0.71107	0.61084	0.55434	0.30279	alk RMSE	Quarterly 1	BVAR	0.87271	0.70078	0.59755	0.51233	0.25608	0.21356	0.10260
SE over Random V	JR/CNY Ratio	UIPUC	1.16467	1.24825	1.31116	1.42947	1.56333	1.41400	1.01501	E over Random W	JR/CNY Ratio	BVARD	0.89569	0.96009	1.01052	1.03940	1.12776	0.99831	0.84158
Parity Model RMS	Quarterly El	UIP	1.21128	1.34308	1.43216	1.54122	1.67310	1.60389	1.23934	utoregression RMS	Quarterly EU	BVAR	1.00885	1.12970	1.20143	1.23066	1.00633	1.13224	0.97317
Uncovered Interest	m R/USD $ m Ratio$	UIPUC	0.89005	0.86706	0.86187	0.79185	0.906532	0.74458	0.88976	Bayesian Vector A	R/USD Ratio	BVARD	0.53257	0.95134	1.08046	1.19427	1.48616	1.62476	1.97836
Panel D.	Quarterly EU	UIP	0.89165	0.82421	0.75723	0.68532	0.74094	0.62529	0.67245	Panel E.	Quarterly EU	BVAR	1.00518	1.11604	1.12628	1.10915	1.28325	1.60066	1.61392
		Time Horizon	1 quarter	2 quarters	3 quarters	4 quarters	6 quarters	8 quarters	12 quarters			Time Horizon	1 quarter	2 quarters	3 quarters	4 quarters	6 quarters	8 quarters	12 quarters

(Continued)
E Ratios
k. RMS
: Anney
Table 2

			Pane	el A. Hoone	<u>er Morton v</u>	s. Hooner	<u> Vorton wit</u> l	Divisia.				
	Ő	uarterly EU	R/USD Ra	tio	Qu	arterly EU	R/CNY Ra	tio	Qu	larterly USI	O/CNY Rat	io
Time Horizon	HM DM	HM	HMD	HMD p-	HM DM	HM	HMD	HMD p-	HM DM	HM	HMD	HMD p-
	stat	p-val	DM stat	val	stat	p-val	DM stat	val	stat	p-val	DM stat	val
1 quarter	1.62088	0.94748	1.68296	0.95381	1.32225	0.90696	1.31475	0.90570	-0.35540	0.36115	-0.31714	0.37557
2 quarters	1.63696	0.94918	1.80845	0.96473	1.48831	0.93167	2.27913	0.98867	-0.15744	0.43745	0.01040	0.50415
3 quarters	1.88984	0.97061	1.95635	0.97479	1.76018	0.96081	2.83148	0.99768	0.02576	0.51027	0.30858	0.62118
4 quarters	1.62204	0.93719	1.73399	0.95854	1.97375	0.96576	3.29031	0.99950	0.07959	0.53093	0.61126	0.72949
6 quarters	1.72854	0.95805	1.89052	0.97066	2.06265	0.98043	3.55787	0.99981	0.98305	0.83721	1.35775	0.91273
8 quarters	1.91496	0.97225	-1.23830	0.10780	2.25430	0.98791	3.35639	0.99961	1.93967	0.97379	3.02105	0.99874
12 quarters	0.20072	0.57954	0.49113	0.68833	2.37015	0.99111	3.55964	0.99981	7.50092	1	73.32533	1
				Panel B.	Sticky Price	e vs. Sticky	Price Divis	sia				
	õ	uarterly EU	R/USD Ra	tio	Qu	arterly EU.	R/CNY Ra	tio	Qu	uarterly USI	O/CNY Rat	io
Time Horizon	SP DM	SP p-val	SPD	SPD p-	SP DM	SP p-val	SPD	SPD p-	SP DM	SP p-val	SPD	SPD p-
	stat		DM stat	val	stat		DM stat	val	stat		DM stat	val
1 quarter	1.58061	0.94302	1.79790	0.96390	0.41866	0.66227	1.25346	0.89498	0.12132	0.54828	-0.12710	0.44943
2 quarters	1.67886	0.95341	2.08308	0.98138	0.04290	0.51711	1.53348	0.93742	0.51641	0.69722	0.21174	0.58384
3 quarters	1.90124	0.97136	2.45306	0.99292	-0.91075	0.18121	0.87378	0.80888	0.39236	0.65260	0.26489	0.60445
4 quarters	1.75032	0.95997	2.58995	0.99520	-1.25270	0.10516	0.32750	0.62835	0.29209	0.61489	0.55272	0.70977
6 quarters	1.89796	0.97115	2.96231	0.99847	-1.22427	0.11042	-0.93538	0.17480	1.12174	0.86901	1.39745	0.91886
8 quarters	1.90176	0.97140	3.87848	0.99995	-1.51155	0.06532	-1.74436	0.04055	2.09082	0.98173	3.23705	0.99940
12 quarters	0.13141	0.55227	0.91818	0.82074	-2.90178	0.00186	-4.79984	0.00001	6.19151	1	5.94435	1
				Panel C. F	lexible Price	e vs. Flexib	le Price Di	visia				
	õ	uarterly EU	R/USD Ra	tio	Qu	arterly EU.	R/CNY Ra	tio	Qu	larterly USI	O/CNY Rat	io
Time Horizon	FP DM	FP p-val	FPD	FPD p-	FP DM	FP p-val	FPD	FPD p-	FP DM	FP p-val	FPD	FPD p-
	stat		DM stat	val	stat		DM stat	val	stat		DM stat	val
1 quarter	1.05825	0.85503	1.28153	0.90000	0.81757	0.79320	1.45818	0.92760	0.41090	0.65943	0.04826	0.51925
2 quarters	0.99477	0.84008	1.64102	0.94960	0.49253	0.68883	1.61354	0.94669	0.82801	0.79617	0.40012	0.65546
3 quarters	0.81382	0.79213	1.66741	0.95228	-0.67501	0.24983	0.84688	0.80147	0.71810	0.76365	0.43634	0.66870
4 quarters	0.61317	0.73012	1.62370	0.94778	-1.15405	0.12424	0.26265	0.60359	0.61410	0.73042	0.67813	0.75115
6 quarters	0.76593	0.77814	1.65757	0.95130	-1.29113	0.09833	-0.77783	0.21834	1.23771	0.89209	1.45537	0.92722
8 quarters	0.65593	0.74406	1.83469	0.966728	-1.94064	0.02615	-1.68113	0.04637	2.18832	0.98568	3.32946	0.99956
12 quarters	-1.29898	0.09698	-0.35665	0.36068	-3.76556	0.00008	-4.57528	0.000002	5.57997	1	5.60950	<u></u>

Table 3: Annex. Diebold-Mariano Test (Full Sample)

(Continued)	
Test	
Diebold-Mariano	
Annex.	
Table 3:	

Pane Pane	ne.	I D. Uncc	overed Inter	rest Parity $\overline{\Omega_{m}}$	vs. Uncover	ed Interest	Parity Use	r Costs	artarly IIC	D/CNV Ref	
ו ג		11/ UNL LIGI	010	ייין אין אין אין אין אין אין אין אין אין		TV/ VIV I LUA	010		Idition the United	D/UNI TIM	01
UIP	þ	UIPUC	UIPUC	UIP DM	UIP p-	UIPUC	UIPUC	UIP DM	UIP p-	UIPUC	UIPUC
val		DM stat	p-val	stat	val	DM stat	p-val	stat	val	DM stat	p-val
0.77	473	-0.05765	0.47701	1.03684	0.85009	0.84633	0.80132	0.98807	0.83844	1.24687	0.89378
0.77	908	-0.55962	0.28787	1.58783	0.94384	1.31472	0.90570	1.32846	0.90799	1.27420	0.89870
0.69	334	-0.79960	0.21197	1.74364	0.95939	1.34434	0.91058	1.73067	0.95824	1.31659	0.90601
0.6	2516	-0.81883	0.20644	1.60640	0.94591	1.12354	0.86940	1.73678	0.95879	1.15737	0.87644
0.6(0827	-0.72203	0.23514	-0.79626	0.21294	-0.60830	0.27149	1.79452	0.96363	1.22200	0.88915
0.59)220	-1.14631	0.12583	-1.48729	0.06847	-1.47577	0.07000	2.37949	0.99133	2.28210	0.98876
0.0	3040	-1.90600	0.02832	-4.28897	0.0001	-4.34558	0.0001	6.33747	1	4.11702	0.999998
	Panel]	E. Bayesian	Vector Au	utoregression	ı vs. Bayesi	ian Vector <i>i</i>	Autoregress	ion Divisia			
arte	rly EUI	R/USD Rat	tio	Qu	larterly EU.	R/CNY Ra	tio	Qu	arterly US	D/CNY Rat	io
BV	AR	BVARD	BVARD	BVAR	BVAR	BVARD	BVARD	BVAR	BVAR	BVARD	BVARD
gv-q	a.l	DM stat	p-val	DM stat	p-val	DM stat	p-val	DM stat	p-val	DM stat	p-val
0.8	9327	-2.18260	0.01453	2.12625	0.98326	-0.94207	0.17308	0.73479	0.76877	0.91306	0.81939
0.9!	5348	-0.95826	0.16897	2.34526	0.99049	-0.54646	0.29238	0.98971	0.83884	0.96308	0.83225
0.9	3278	-0.38488	0.35016	2.07342	0.98093	-0.54856	0.29165	0.96801	0.83348	0.84410	0.80069
0.65	5851	0.70221	0.75873	1.93035	0.97322	-0.41061	0.34068	-0.07746	0.46913	-0.10947	0.45641
0.0	3527	0.93571	0.82529	-2.01742	0.02183	0.02925	0.51167	-1.86141	0.03134	-1.40051	0.08068
0.6'	7399	0.78989	0.78520	-0.22562	0.41075	0.51576	0.69699	0.44809	0.67295	1.02530	0.84739
0.65	812	0.41820	0.66210	1.27720	0.89923	0.79214	0.78586	-2.50060	0.00620	4.81887	1.00000

	1	1										<u> </u>	1																				
	io	HMD p-	val	0.22176	0.31488	0.47006	0.62095	0.84887	0.97126	0.94590		io	SPD p-	val	0.01269	0.00110	0.00018	0.000004	0	0	0			io	FPD p-	val	0.01511	0.00133	0.00043	0.00002	0	0	0
	J/CNY Ra	HMD	DM stat	-0.76626	-0.48205	-0.07511	0.30798	1.03158	1.89965	1.60633		D/CNY Ra	SPD	DM stat	-2.23549	-3.06090	-3.57131	-4.46768	-6.42924	-8.35566	-9.74436			J/CNY Ra	FPD	DM stat	-2.16718	-3.00533	-3.33408	-4.16834	-5.74837	-7.71146	-8.00726
	arterly USI	HM	p-val	0.20422	0.26099	0.33035	0.39789	0.70761	0.91779	0.99365		arterly USI	SP p-val		0.00974	0.00073	0.00012	0.000002	0	0	0			arterly USI	FP p-val		0.01531	0.00115	0.00014	0.000002	0	0	0
	Qu	HM DM	stat	-0.82663	-0.64031	-0.43894	-0.26092	0.54641	1.39033	2.49233		Qu	SP DM	stat	-2.33632	-3.18130	-3.67710	-4.61095	-6.42245	-7.91021	I	11.38533		Qu	FP DM	stat	-2.16199	-3.04877	-3.62995	-4.58597	-6.21818	-7.71327	-9.63686
ı Divisia	tio	HMD p-	val	0.90570	0.98867	0.99768	0.99950	0.99981	0.99961	0.99981	sia	tio	SPD p-	val	0.97224	0.99769	0.99945	0.99995	0.99999	0.99990	1.00000		risia	tio	FPD p-	val	0.96503	0.99827	0.99971	0.999999	1.00000	0.99991	1.00000
Morton with	R/CNY Ra	HMD	DM stat	1.31475	2.27913	2.83148	3.29031	3.55787	3.35639	3.55964	Price Divis	R/CNY Ra	SPD	DM stat	1.91478	2.83261	3.26569	3.87530	4.19430	3.72703	4.82948		le Price Div	R/CNY Ra	FPD	DM stat	1.81235	2.92315	3.44367	4.27723	4.47460	3.75127	4.50312
s. Hooper l	uarterly EU	HM	p-va.l	0.90696	0.93167	0.96081	0.96576	0.98043	0.98791	0.99111	e vs. Sticky	arterly EU	SP p-val		0.95650	0.99275	0.99443	0.99811	0.99992	0.99997	0.99996		e vs. Flexib	arterly EU	FP p-val		0.95671	0.99563	0.99732	0.99958	1.00000	1.00000	0.99994
er Morton v	0r	HM DM	stat	1.32225	1.48831	1.76018	1.97375	2.06265	2.25430	2.37015	Sticky Price	0 0	SP DM	stat	1.71144	2.44476	2.53854	2.89649	3.78456	4.00098	3.94176		lexible Price	Q	FP DM	stat	1.71373	2.62169	2.78481	3.34101	4.57129	4.49366	3.83099
el A. Hoope	tio	HMD p-	va.l	0.49762	0.68223	0.75220	0.70194	0.99781	0.99856	0.99905	Panel B.	tio	SPD p-	val	0.67737	0.86788	0.89003	0.86745	0.99973	0.99988	0.99842		Panel C. F	tio	FPD p-	val	0.77367	0.94360	0.96568	0.97598	0.99998	0.99998	0.99997
Pan	R/USD Ra	HMD	DM stat	-0.00596	0.47394	0.68143	0.52998	2.85006	2.97941	3.10483		R/USD Ra	SPD	DM stat	0.46036	1.11643	1.22670	1.11442	3.45535	3.66973	2.95165			R/USD Ra	FPD	DM stat	0.75098	1.58569	1.82076	1.97699	4.14073	4.15835	4.03818
	uarterly EU	HM	p-val	0.14663	0.00967	0.02424	0.03160	0.03831	0.06766	0.99875		uarterly EU	SP p-val		0.16732	0.01028	0.04494	0.03383	0.04480	0.17676	0.94252			uarterly EU	FP p-val		0.17822	0.02107	0.03420	0.02967	0.02754	0.27207	0.93908
	Ō	HM DM	stat	-1.05101	-2.33878	-1.97318	-2.01785	-1.77066	-1.49347	3.02341		õ	SP DM	stat	-0.96480	-2.31604	-1.69604	-1.82733	-1.69755	-0.92780	1.57628			Q	FP DM	stat	-0.92218	-2.03222	-1.82237	-1.88567	-1.91818	-0.60657	1.54710
		Time Horizon		1 quarter	2 quarters	3 quarters	4 quarters	6 quarters	8 quarters	12 quarters			Time Horizon		1 quarter	2 quarter	3 quarter	4 quarter	6 quarter	8 quarter	12 quarter				Time Horizon		1 quarter	2 quarter	3 quarter	4 quarter	6 quarter	8 quarter	12 quarter

Table 4: Annex. Diebold-Mariano Test (W/out Negative Rates)

Table 4: Annex. Diebold-Mariano Test (Continued)

Osts	Quarterly USD/CNY Ratio	IP DM UIP P- UIPUC UIPUC	at val DM stat p-val	0.46802 0.31989 -1.19085 0.11686	0.77960 0.21781 -2.04567 0.02039	.04770 0.14739 -2.80059 0.00255	.53196 0.06277 -3.80415 0.00007	0.55325 0.00534 -5.77617 0	3.38148 0.00036 -7.20690 0		11.75377	Divisia	Quarterly USD/CNY Ratio	VAR BVAR BVARD BVARD	M stat p-val DM stat p-val	0.73967 0.22975 -1.50778 0.06581	.79010 0.03672 -2.68014 0.00368	0.46537 0.00684 -3.14333 0.00084	0.83998 0.00226 -3.30003 0.00048 0.00048	0.00280 0 -6.85450 0	7.58636 0 -7.66896 0	0 - 0	10.75305
Parity U	atio	UIPUC	p-val	0.97290	0.99161	0.99829	0.999990	0.999999	0.99504	0.53343		Autoregre	atio	BVARL	p-val	0.07663	0.22692	0.55761	0.66764	0.95748	0.49146	0.01733	
ed Interest	R/CNY R	UIPUC	DM stat	1.92516	2.39167	2.92648	3.71277	4.17411	2.57844	0.08388		ian Vector	R/CNY Ra	BVARD	DM stat	-1.42809	-0.74903	0.14492	0.43340	1.72215	-0.02141	-2.11222	
/s. Uncover	larterly EU	UIP p-	val	0.98333	0.99990	1.00000	1	1	1.00000	0.99506		1 vs. Bayes:	larterly EU	BVAR	p-val	0.56505	0.96070	0.96339	0.97884	0.56359	0.95069	0.38487	
est Parity v	Qu	UIP DM	stat	2.12791	3.72558	4.74068	5.64407	5.95363	5.00673	2.57971		toregression	Qu	BVAR	DM stat	0.16378	1.75885	1.79143	2.03042	0.16009	1.65162	-0.29272	
vered Inter	tio	UIPUC	p-val	0.01629	0.01379	0.05148	0.05243	0.23575	0.01782	0.06537		Vector Au	tio	BVARD	p-val	0.00212	0.27405	0.77799	0.94060	0.99448	0.99900	0.99364	
nel D. Unco	R/USD Rai	UIPUC	DM stat	-2.13715	-2.20313	-1.63063	-1.62171	-0.72006	-2.10111	-1.51122		E. Bayesian	R/USD Rai	BVARD	DM stat	-2.85913	-0.60060	0.76543	1.55985	2.54159	3.09159	2.49148	
Pa	uarterly EU	UIP p-	val	0.01069	0.00204	0.00718	0.00764	0.03496	0.00558	0.02598		Panel	uarterly EU	BVAR	p-val	0.52857	0.88552	0.83187	0.81430	0.95497	0.99751	0.98912	
	dr dr	UIP DM	stat	-2.30108	-2.87234	-2.44808	-2.42570	-1.81245	-2.53764	-1.94346			Q	BVAR	DM stat	0.07168	1.20306	0.96159	0.89385	1.69506	2.80838	2.29463	
		Time Horizon		1 quarter	2 quarter	3 quarter	4 quarter	6 quarter	8 quarter	12 quarter	1			Time Horizon		1 quarter	2 quarter	3 quarter	4 quarter	6 quarter	8 quarter	12 quarter	

9 Annex: Actual vs Forecasted values, for 3 months, 6 months 9 months and 12 months ahead





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